April 11th, 2011

## NAME:

Total number of Marks: /25

Answer all the questions (questions are on <u>both</u> sides of the paper).

Marks will be given for clarity of the solution.

Write your solutions in the question paper.

Good Luck!

(1) Suppose that Z is a Weibull random variable with density  $f(x; \phi, \alpha) = \left(\frac{\alpha}{\phi}\right) \left(\frac{x}{\phi}\right)^{\alpha-1} \exp(-(x/\phi)^{\alpha})$ . Show that

$$\mathbb{E}(Z^r) = \phi^r \Gamma\left(1 + \frac{r}{\alpha}\right).$$
[3]

Hint: Use

$$\int x^a \exp(-x^b) dx = \frac{1}{b} \Gamma(\frac{a}{b} + \frac{1}{b}) \qquad a, b > 0.$$

This result may be useful in the questions below.

(2) Let us suppose that the random variable X is a mixture of Weibull distributions

$$f(x;\theta) = p(\frac{\alpha_1}{\phi_1})(\frac{x}{\phi_1})^{\alpha_1 - 1} \exp(-(x/\phi_1)^{\alpha_1}) + (1-p)(\frac{\alpha_2}{\phi_2})(\frac{x}{\phi_2})^{\alpha_2 - 1} \exp(-(x/\phi_2)^{\alpha_2}).$$

[3]

(i) Derive the mean and variance of X.

(ii) Obtain the exponential distribution which best fits the above mixture Weibull according to the Kullbach-Lieber criterion (recall that the exponential is  $g(x; \lambda) = \frac{1}{\lambda} \exp(-x/\lambda)$ ). [3]

- (3) Let us suppose that  $\{T_i\}_i$  are the survival times of lightbulbs. We will assume that  $\{T_i\}$  are iid random variables with the density  $f(\cdot; \theta_0)$  and survival function  $\mathcal{F}(\cdot; \theta_0)$ , where  $\theta_0$  is unknown. The survival times are censored, and  $Y_i = \min(T_i, c)$  and  $\delta_i$  are observed (c > 0), where  $\delta_i = 1$  if  $Y_i = T_i$  and is zero otherwise.
  - (a) (i) State the log-likelihood of  $\{(Y_i, \delta_i)\}_i$ .

(ii) We denote the above log-likelihood as  $\mathcal{L}_T(\theta)$ . Show that

$$-\mathbb{E}\left(\frac{\partial^{2}\mathcal{L}_{T}(\theta)}{\partial\theta^{2}}\rfloor_{\theta=\theta_{0}}\right)=\mathbb{E}\left(\frac{\partial\mathcal{L}_{T}(\theta)}{\partial\theta}\rfloor_{\theta=\theta_{0}}\right)^{2},$$

stating any important assumptions that you may use.

[3]

[1]

- (b) Let us suppose that the above survival times satisfy a Weibull distribution  $f(x; \phi, \alpha) = (\frac{\alpha}{\phi})(\frac{x}{\phi})^{\alpha-1} \exp(-(x/\phi)^{\alpha})$  and as in part (a) we observe and  $Y_i = \min(T_i, c)$  and  $\delta_i$ , where c > 0.
  - (i) Using your answer in part 2a(i), give the log-likelihood of  $\{(Y_i, \delta_i)\}_i$  for this particular distribution (we denote this as  $\mathcal{L}_T(\alpha, \phi)$ ) and derive the profile likelihood of  $\alpha$  (profile out the nusiance parameter  $\phi$ ).

Suppose you wish to test  $H_0$ :  $\alpha = 1$  against  $H_A$ :  $\alpha \neq 1$  using the loglikelihood ratio test, what is the limiting distribution of the test statistic under the null? [3] (ii) Let  $\hat{\phi}_T$ ,  $\hat{\alpha}_T = \arg \max \mathcal{L}_T(\alpha, \phi)$  (maximum likelihood estimators involving the censored likelihood). Do the estimators  $\hat{\phi}_T$  and  $\hat{\alpha}_T$  converge to the true parameters  $\phi$  and  $\alpha$  (you can assume that  $\hat{\phi}_T$  and  $\hat{\alpha}_T$  converge to some parameters, and your objective is to find whether these parameters are  $\phi$  and  $\alpha$ ). [3]

(iii) Obtain the (expected) Fisher information matrix of maximum likelihood estimators.

(iv) Using your answer in part 2b(iii) derive the limiting variance of the maximum likelihood estimator of  $\hat{\alpha}_T$ . [3]