

Multiple-try MCMC without Rejection

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In this note, I explain how the famous multiple-try Metropolis (MTM) algorithm [5] can be turned into a rejection-free MCMC method without extra computational cost.

For details, see our paper [arxiv] (see Algorithm 7 therein) or my other slides at <https://web.stat.tamu.edu/~quan/>.

MTM algorithm

- ▶ \mathcal{X} : a general state space.
- ▶ $Q(x, \cdot)$: a proposal distribution given current state $x \in \mathcal{X}$.
- ▶ $q(x, \cdot)$: density function of $Q(x, \cdot)$; we assume $q(x, y) = q(y, x)$.
- ▶ π : density function of the target distribution.

MTM is essentially a Metropolis-Hastings algorithm with a complicated proposal scheme. Instead of simply proposing one state from $Q(x, \cdot)$, MTM proposes multiple candidate moves (i.e., multiple "tries") and then assign larger proposal probabilities to states with larger π . We assume the weight of y given current state x is proportional to

$$h\left(\frac{\pi(y)}{\pi(x)}\right)$$

for some function $h: (0, \infty) \rightarrow (0, \infty)$.

MTM algorithm

An iteration of MTM at state x with m tries:

1. Draw y_1, \dots, y_m from $Q(x, \cdot)$.
2. Select y from y_1, \dots, y_m with probability $\propto h\left(\frac{\pi(y)}{\pi(x)}\right)$.
3. Draw x_1, \dots, x_{m-1} from $Q(y, \cdot)$. Set $x_m = x$.
4. Accept y with probability

$$\min \left\{ 1, \frac{Z_h(x, y_1, \dots, y_m)}{Z_h(y, x_1, \dots, x_m)} \right\},$$

where $Z_h(x, y_1, \dots, y_m) = \sum_{k=1}^m h\left(\frac{\pi(y_k)}{\pi(x)}\right)$.

Choice of h

The recent studies [2, 1] suggest that one wants to choose h such that

$$h(u) = u h(u^{-1}), \quad \forall u > 0.$$

Such a function is called a balancing function. Examples include

$$h(u) = 1 + u, \quad h(u) = \sqrt{u}, \quad h(u) = \min\{1, u\}.$$

We assume h is a balancing function henceforth.

Our Multiple Try Importance Tempering algorithm

An iteration of MT-IT at state x with m tries, y_1, \dots, y_m .

1. Select y from y_1, \dots, y_m with probability $\propto h \left(\frac{\pi(y)}{\pi(x)} \right)$.
2. Accept y . Assign to the previous state x (un-normalized) importance weight $Z_h(x, y_1, \dots, y_m)^{-1}$.
3. Draw x_1, \dots, x_{m-1} from $Q(y, \cdot)$. Set $x_m = x$. In the next iteration, we use x_1, \dots, x_m as the m tries at state y .

The *only differences* from MTM are that (1) we always accept y , (2) we calculate importance weight instead of acceptance probability (note both rely on evaluating the function Z_h).

Why is it correct?

Let's consider the dynamics of the state $(x, \{y_1, \dots, y_m\}) \in \mathcal{X} \times \mathcal{X}_m$, where \mathcal{X}_m is the collection of all *unordered* subsets of \mathcal{X} with m elements. It is a Markov chain with transition density

$$p((x, \{y_1, \dots, y_m\}), (y, \{x_1, \dots, x_m\})) = A_1 A_2,$$

where

$$A_1 = \frac{h\left(\frac{\pi(y)}{\pi(x)}\right)}{Z_h(x, y_1, \dots, y_m)}, \quad A_2 = \prod_{k=1}^{m-1} q(y, x_k).$$

A_1 corresponds to how we select y from $\{y_1, \dots, y_m\}$, and A_2 corresponds to how we generate x_1, \dots, x_m . *Note that x_m is fixed to be x !*

Why is it correct?

p satisfies the detailed balance condition w.r.t. the stationary distribution

$$\pi_h(x, \{y_1, \dots, y_m\}) \propto \pi(x) Z_h(x, y_1, \dots, y_m) \prod_{k=1}^m q(x, y_k).$$

Comparing this to a reference distribution

$$\bar{\pi}(x, \{y_1, \dots, y_m\}) = \pi(x) \prod_{k=1}^m q(x, y_k),$$

we see that $Z_h(x, y_1, \dots, y_m)^{-1}$ is the importance weight we need.

Why is it correct?

Here is the proof of the detailed balance condition. Since (y_1, \dots, y_m) is treated as an unordered set, we can assume $y = y_k$. Then,

$$\begin{aligned} & \pi_h(x, \{y_1, \dots, y_m\})p((x, \{y_1, \dots, y_m\}), (y, \{x_1, \dots, x_m\})) \\ &= \pi(x)q(x, y)h\left(\frac{\pi(y)}{\pi(x)}\right) \prod_{k=1}^{m-1} q(y, x_k) \prod_{k=1}^{m-1} q(x, y_k). \end{aligned}$$

It only remains to show that

$$\pi(x)q(x, y)h\left(\frac{\pi(y)}{\pi(x)}\right) = \pi(y)q(y, x)h\left(\frac{\pi(x)}{\pi(y)}\right),$$

which holds since q is assumed symmetric and h is a balancing function.

Why is it correct?

Caveat: Fixing $x_m = x$ is important! This guarantees that the transition density from $(x, \{y_1, \dots, y_m\})$ to $(y, \{x_1, \dots, x_m\})$ is nonzero only if

$$y \in \{y_1, \dots, y_m\} \text{ and } x \in \{x_1, \dots, x_m\}.$$

Without this symmetry, reversibility fails and the stationary distribution of the chain is unclear.

References

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