

Lecture 1: March 28

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Theorem 1.1. *Fix $\alpha \in (0, 1)$. Then for any $\epsilon \in (0, 1)$,*

$$\int \left\{ \frac{1}{n} D_\alpha^{(n)}(\theta, \theta_0) \right\} \Pi_{n,\alpha}(d\theta | X^{(n)}) \leq \frac{\alpha}{n(1-\alpha)} \int r_n(\theta, \theta_0) \rho(d\theta) + \frac{1}{n(1-\alpha)} D(\rho, \Pi_n) + \frac{1}{n(1-\alpha)} \log(1/\epsilon)$$

\forall probability measure $\rho \ll \Pi$, with $\mathbb{P}_{\theta_0}^{(n)}$ probability at least $(1 - \epsilon)$.

Let $T_1 = \int r_n(\theta, \theta_0) \rho(d\theta)$ and $T_2 = D(\rho | \Pi_n)$.

Comments

1. There is a trade off going on in between T_1 and T_2 . We need ρ to place more and more mass near θ^* to make T_1 small. We have to set $\rho = \Pi_n$ to make T_2 small. Our choice of ρ [ρ : the prior Π_n restricted to $B_n(\theta^*, \epsilon; \theta_0)$] provides an optimal trade off.
2. Let $\psi(\rho) = T_1 + T_2$. $\psi(\rho)$ is minimized (over all $\rho \ll \Pi_n$) when $\rho = \Pi_{n,\alpha}$ and $\psi(\Pi_{n,\alpha}) = -\log$ marginal likelihood under the fractional posterior.
3. The eventual bound (after you choose ρ as mentioned) only depends on the ‘‘prior concentration’’ ($\Pi_n[B_n(\theta^*, \epsilon; \theta_0)]$). You do not need the metric entropy and testing conditions to obtain this risk bounds. Also unified analysis of well-specified and miss-specified models can be obtained. This in particular allows to work with heavy-tailed priors.