

April 5, 2018

Name:

STAT 632 (Anirban Bhattacharya)

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The exam is out of 60 points. For each problem a maximum of 20 points can be earned. You have 65 minutes to complete the exam.

- Work your problems in the space provided.
- Show all work clearly and concisely.
- Justify your answers. State results you use.
- Box your final answer whenever applicable.
- $\mathbb{1}_A(x)$ denotes the indicator function of the set A , that is, $\mathbb{1}_A(x) = 1$ if $x \in A$ and is zero otherwise.

Distribution name	pdf/pmf	mean	variance
Binomial(n, p)	$\binom{n}{x} p^x (1-p)^{n-x}, x \in \{0, 1, \dots, n\}$	np	$np(1-p)$
Poisson(λ)	$\frac{e^{-\lambda} \lambda^x}{x!}, x \in \{0, 1, \dots\}$	λ	λ
Gamma(α, β)	$\frac{\beta^\alpha}{\Gamma(\alpha)} e^{-\beta x} x^{\alpha-1}, x \in (0, \infty)$	α/β	α/β^2
Beta(α, β)	$\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}, x \in (0, 1)$	$\alpha/(\alpha+\beta)$	
$N(\mu, \sigma^2)$	$\frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, x \in \mathbb{R}$	μ	σ^2
$N_d(\mu, \Sigma)$	$(2\pi)^{-d/2} \Sigma ^{-1/2} e^{-\frac{1}{2}(x-\mu)'\Sigma^{-1}(x-\mu)}, x \in \mathbb{R}^d$	μ	Σ

1. Suppose x_1, \dots, x_n are independent observations drawn from a $N(\lambda_1\eta, 1)$ distribution, and y_1, \dots, y_m are independent observations drawn from a $N(\lambda_2\eta, 1)$ distribution. Further, assume that $\mathbf{x} = (x_1, \dots, x_n)$ and $\mathbf{y} = (y_1, \dots, y_m)$ are conditionally independent given $\lambda_1, \lambda_2, \eta$. Assume independent $N(0, 1)$ priors on $\lambda_1, \lambda_2, \eta$.

(a) Clearly outline the steps of a **Gibbs sampler** to sample from the posterior distribution of $(\lambda_1, \lambda_2, \eta) \mid (\mathbf{x}, \mathbf{y})$, specifying the necessary **full conditional distributions**.

(b) Suppose you run the Gibbs sampler for 11,000 iterations and discard the first 1,000 iterations as burn-in. Describe how you shall use the output of the Gibbs sampler to obtain a point estimate and 95% confidence interval for $\lambda_1\lambda_2$.

2. Suppose we observe $y = (y_1, \dots, y_n)$, with $y_i \mid \mu_i \sim N(\mu_i, 1)$ independently for $i = 1, \dots, n$. Given $\omega \in (0, 1)$, assume independent $N(0, (1 - \omega)/\omega)$ priors on the μ_i s. [**Note:** $(1 - \omega)/\omega$ is the prior variance, which is guaranteed to be positive since $\omega \in (0, 1]$]

Assume a $\text{Beta}(1/2, 1/2)$ prior on ω .

(a) Show that the marginal posterior distribution of $\omega \mid y$ is of the form

$$\pi(\omega \mid y) \propto \omega^{A-1} (1 - \omega)^{B-1} e^{-C\omega}, \quad \omega \in (0, 1).$$

Identify A, B, C .

(b) Describe the steps of a Metropolis algorithm to sample from $\pi(\omega \mid y)$ that performs a normal random walk (with standard deviation σ_{MH}) on $\log\left(\frac{\omega}{1-\omega}\right)$. In particular, your answer should contain an explicit expression for the acceptance probability.

(c) Describe the steps of a Bootstrap filter to obtain a discrete approximation to $\pi(\omega \mid y)$. How shall you use this to estimate $P(\omega > 0.5 \mid y)$?

3. For p and q two probability densities, recall the total variation distance

$$d_{TV}(p, q) = \frac{1}{2} \int |p - q| = \int_{p > q} (p - q) = \int_{q > p} (q - p),$$

Let $p(x) = e^{-x} \mathbb{1}_{(0, \infty)}(x)$ and $q(x) = \lambda e^{-\lambda x} \mathbb{1}_{(0, \infty)}(x)$ be the densities of exponential distributions with rates 1 and $\lambda > 1$ respectively.

(a) Find $d_{TV}(p, q)$.

(b) Let q_n be the density of an exponential distribution with rate $\lambda_n = (1 + 1/n)$. Show that $\lim_{n \rightarrow \infty} d_{TV}(p, q_n) = 0$. [**Hint:** You may need to use the fact $\lim_{n \rightarrow \infty} (1 + 1/n)^n = e$]

